

A COMPUTATIONAL TECHNIQUE FOR THE
ANALYSIS OF TWO-WAY CLASSIFICATION
WITH DISPROPORTIONATE SUBCLASS FREQUENCIES

By

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Introduction

The analysis of a two-way classification with inter-action is sometimes not possible even with the use of computers because of the size of the matrix involved. The purpose of this paper is to partition the matrix and work with submatrices of manageable sizes that most often the calculations can be done with a desk calculator. The method can also be used using a computer when the ordinary least squares analysis cannot be used.

The method evolved in this paper uses the least squares procedure. It is an exact and convenient method for the randomized complete block design with missing values provided the rank of the design matrix X is equal to no. of blocks + no. of treatments - 1. It can also be used to analyze a factorial experiment in CRD with two factors and unequal number of replications. All balanced incomplete block designs not utilizing inter block information can also be analyzed using this method and tests for interaction is also possible with this method. Built-in checks are incorporated in the procedure so that at certain stages the computation can be checked.

Consider the model

$$(1) Y = X\beta + \epsilon$$

where Y is an $n \times 1$ vector of observations

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ϵ is an $n \times 1$ random vector

X is an $n \times (a \neq b)$ matrix of known fixed quantities

β is a $(a \neq b) \times 1$ vector of unknown parameters

Assume that $\epsilon \sim N(0, \sigma^2 I)$ and the data is such that the rank of X is $(a \neq b) - 1$.

$$\text{Let } \beta = \begin{Bmatrix} a \\ a \times 1 \\ \rho \\ b \times 1 \end{Bmatrix}$$

Data: ρ

	1	2	...	b	Total
1	Y_{111}	Y_{121}	.	Y_{1b1}	$Y_{1..}$
	Y_{112}	Y_{122}	.	Y_{1b2}	
	\vdots	\vdots	.	\vdots	
	$Y_{11n_{11}}$	$Y_{12n_{12}}$.	$Y_{1bn_{1b}}$	
a	Y_{211}	.			$Y_{2.}$
	Y_{212}	.			
	\vdots	.			
	$Y_{21n_{21}}$.			
a		Y_{a21}		Y_{ab1}	$Y_{a..}$
		Y_{a22}		Y_{ab2}	
		\vdots		Y_{abn}	
		$Y_{a2n_{a2}}$		$Y_{abn_{ab}}$	
Total	Y.1.	Y.2.	...	Y.b.	

No. of Observations

ρ

	1	2	j	b	Total
1	n_{11}	n_{12}	n_{1j}	n_{1b}	$n_{1.}$
2	n_{21}	n_{22}	n_{2j}	π	$n_{2.}$
a i			n_{ij}		$n_{i.}$
a	n_{a1}	n_{a2}	n_{aj}	n_{ab}	$n_{a..}$
Total	$n_{.1}$	$n_{.2}$	$n_{.j}$	$n_{.b}$	$n_{..}$

NOTATIONS

Let

$$A = \begin{Bmatrix} Y_{1..} \\ Y_{2..} \\ \vdots \\ Y_{a..} \end{Bmatrix} \quad B = \begin{Bmatrix} Y_{.1} \\ Y_{.2} \\ \vdots \\ Y_{.b} \end{Bmatrix}$$

$$N = \begin{matrix} a \times b \\ \left\{ n_{ij} \right\} \end{matrix}$$

$$D_a = \begin{matrix} a \times a \\ \left\{ \begin{matrix} n_{2.} & 0 & \dots & 0 \\ 0 & n_{2.} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots \\ 0 & \dots & 0 & \dots & n_{a.} \end{matrix} \right\} \end{matrix}$$

$$\text{and } \begin{matrix} D_b \\ b \times b \end{matrix} = \left\{ \begin{array}{ccc} n_{1.} & 0 \dots 0 \\ 0 & n_{.2} \dots 0 \\ \vdots & \ddots & \vdots \\ \dots & \dots & \dots \\ 0 \dots 0 \dots n_{.b} \end{array} \right\}$$

To get P (adjusted for a) form

$$(2) \quad \begin{array}{c|c|c} D_a & N & A \\ \hline & D_b & B \end{array}$$

and to get a (adjusted for P) form

$$(3) \quad \begin{array}{c|c|c} D_b & N' & B \\ \hline & D_a & A \end{array}$$

where N' is the transpose of N .

Let us consider (2). Now partition (2) into I and II as follows:

$$\begin{array}{c|c|c} D_a & N & A \\ \hline & D_b & B \end{array} \left. \begin{array}{l} \} \\ \} \end{array} \right\} \begin{array}{l} \text{I} \\ \text{II} \end{array}$$

We form the following AOVs:

Analysis of Variance
 ρ (adj for α)

Sources of variation	d.f.	S.S.
Total	$n - 1$	$\sum_i \sum_j \sum_k Y_{ijk}^2 - CF$
α (unadj)	$a - 1$	$\sum_i Y_{i...}^2 - CF$
ρ (adj for α)	$b - 1$	$\sum CPID$
$\alpha\rho$	$(a-1)(b-1) - e$	by subtraction
Within Cells	$n - ab + e$	$\sum_i \sum_j \sum_k Y_{ijk}^2 - \sum_i \sum_j \frac{Y_{ij}^2}{n_{ij}}$

where e is the number of empty cells

$$CF \text{ if } \frac{(\sum_i \sum_j \sum_k Y_{ijk})^2}{n..}$$

and $n = n..$

Analysis of Variance
 α (adj for ρ)

Sources of Variation	d.f.	S. S.
Total	$n - 1$	same as above
ρ (unadj)	$b - 1$	$\sum_j \frac{Y^2_{.j}}{n_{.j}} - CF$
α (adj for ρ)	$a - 1$	$\alpha \text{ SS}(\text{unadj}) + \rho \text{ SS}(\text{adj}) - \rho \text{ SS}(\text{unadj})$

$a\beta$ $(-1) (b-1) - e$ Same as above

Within cells $n - ab + e$ Same as above

If $n_{ij} < 2$ for all i and j the within cells cannot be obtained and hence will not appear in the analysis.

If there is no interaction and (adj) MS is significant the following hypothesis maybe tested using a t-test:

$$H_0: \rho_1 - \rho_2 = 0$$

$$\text{Test Statistics: } t = \frac{b_1 - b_2}{S_{(b_1-b_2)}}$$

where $S_{(b_1-b_2)} = \sqrt{EMS (d'_{11} + d'_{22} - 2 d'_{12})}$

and d'_{ij} are elements of D' ; EMS is error mean square

EXAMPLE 1 (Two-Way Classification without interaction)

Randomized complete block with missing observations:

Data:

	ρ (Treatments)						No. of observation				
	1	2	3	4	5		1	2	3	4	5
a	1	5	1	2	8	a	1	1	1	1	3
	2	4	-	3	7		2	1	0	1	2
	3	6	2	1	9		3	1	1	1	3
	4	-	3	2	5		4	0	1	1	2
	5	5	2	3	10		5	1	1	1	3
	20	8	11	39		4	4	5	13		

Form $D_a \mid N \mid A$
 $\mid D_b \mid B$:

3	0	0	0	0	1	1	1	8	}	I
	2	0	0	0	1	0	1	7		
		3	0	0	1	1	1	9		
			2	0	0	1	1	5		
				3	1	1	1	10		
					4	0	0	20	}	II
					4	0	8			
					5	11				

Doolittle I, we get:

3	0	0	0	0	1	1	1	8
1	0	0	0	0	1/3	1/3	1/3	8/3
	2	0	0	0	1	0	1	7
	1	0	0	0	1/2	0	1/2	7/2
	3	0	0		1	1	1	9
	1	0	0		1/3	1/3	1/3	9/3
	2	0			0	1	1	5
	1	0			0	1/2	1/2	5/2
	3				1	1	1	10
	1				1/3	1/3	1/3	10/3

Adjusting II using this result we get

$$c_{11} = 4 - [1(1/3) + 1(1/2) + 1(1/3) + 0(0)] =$$

$$4 - \frac{2+3+2+2}{6} = 4 - \frac{9}{6} = \frac{15}{6}$$

$$c_{12} = - [1(1/3) + 1(0) + 1(1/3) + 0(1/2) + 1(1/3)]$$

$$= - \frac{6}{6}$$

$$c_{13} = - [1(1/3) + 1(1/2) + 1(1/3) + 0(1/2) + 1(1/3)]$$

$$= - \frac{2+3+2+2}{6} = - \frac{9}{6}$$

$$c_{22} = 4 - [1(1/3) + 0(0) + 1(1/3) + 1(1/2) + 1(1/3)]$$

$$= 4 - \frac{9}{6} = \frac{15}{6}$$

$$c_{23} = - [1(1/3) + 0(1/2) + 1(1/3) + 1(1/2) + 1(1/3)]$$

$$= - \frac{9}{6}$$

$$c_{33} = 5 - [1(1/3) + 1(1/2) + 1(1/3) + 1(1/2) + 1(1/3)]$$

$$= 5 - 2 = \frac{18}{6}$$

$$g_1 = 20 - [1(8/3) + 1(7/2) + 1(9/3) + 0(5/2) + (10/3)]$$

$$= 20 - 9 - \frac{7}{2} = \frac{15}{2}$$

$$\begin{aligned}
 g_2 &= 8 - [1(8/3) + 0(7/2) + 1(9/3) + 1(5/2) + 1(10/3)] \\
 &= 8 - \frac{27}{3} - \frac{5}{2} = -\frac{7}{2}
 \end{aligned}$$

$$\begin{aligned}
 g_3 &= 11 - [1(8/3) + 1(7/2) + 1(9/3) + 1(5/2) + 1(10/3)] \\
 &= 11 - 9 - 6 = -4
 \end{aligned}$$

$$\begin{array}{r|l}
 15/6 - 6/6 - 9/6 & 15/2 \\
 15/6 - 9/6 & -7/2 \\
 18/6 & -8/2
 \end{array}$$

add $1/3$ to each element of the LHS of the above, convert to decimals and Doolittle we get:

2.83333333	-.66666667	-1.16666667	7.5	1	0	0
	2.83333333	-1.16666667	-3.5	0	1	0
		3.33333334	-4.0	0	0	1
2.83333333	-.66666667	-1.16666667	7.5	1.10	0	0
1.0	-.23529412	-.41176471	2.64705883	.35294118	0	0
	2.67647058	-1.44117648	-1.73529411	.23529412	1	0
	1.0	-.53846154	-.64835164	.08791208	.37362637	0
		2.07692307	-1.84615382	.53846153	.53846154	1
		1.0	-.88888889	.25925926	.25925926	.48148148
				.51322751	.22751322	.25925925
					.51322751	.25925925
						.48148148

Subtracting 1/3 from the inverse above and denoting as D', we get:

$$D' = \begin{Bmatrix} .1798418 & -.10582011 & -.07407408 \\ & .17989418 & -.07407408 \\ & & .14814815 \end{Bmatrix}$$

Estimates of the Treatment Effects are as follows:

$$b_3 = - .88888889$$

$$b_2 = - .88888889(.53846154) + (-.64835164) (1.0)$$

$$= - 1.12698412$$

$$b_1 = - .88888889(.53846154) + (-.64835164) \\ (.23529412) + (2.64705883) (1.0)$$

$$= 2.01587302$$

$$SS \text{ Treat}(\text{adj for Blocks}) = \sum CPID$$

$$= 7.5(2.64705883) + (-1.73529411)(-.64835164)$$

$$+ (-1.84615383)(-.88888889)$$

$$= 22.61904762$$

AOV

Source	d.f.	S.S.	M.S.	F
Total	12	30.00	—	
Block	4	1.67	—	
Block(adj)	4	1.09	.27	< 1
Treat (adj)	2	22.62	— 11.31	12.26
Treat	2	23.20	—	
Error	6	5.71	.95	

$$\text{Treat SS} = \frac{20^2}{4} + \frac{8^2}{4} + \frac{11^2}{5} - \frac{39^2}{13} = 23.20$$

$$\text{Block}(\text{adj}) \text{ SS} = \text{Block SS} + \text{Treat}(\text{adj}) \text{ SS} - \text{Treat SS} \\ = 24.29 - 23.20 = 1.09$$

Since there is a strong evidence of treatment differences we may test the following hypothesis using the elements of D' and a t-test:

$$H_0: \rho_1 - \rho_3 = 0$$

$$t = \frac{b_1 - b_3}{S_{b_1 - b_3}} = \frac{2.01587302 + .88888889}{.67252} = 4.319$$

$$\begin{aligned} \text{where } S_{b_1 - b_3} &= \sqrt{\text{EMS} \{d'_{11} + d'_{33} - 2d'_{13}\}} \\ &= \sqrt{.95[.17989418 + .14814815 - 2(-.074074080)]} \\ &= \sqrt{.4523809655} \\ &= .67252 \end{aligned}$$

EXAMPLE 2 (Two-Way Classification with interaction)

Date:	ρ				Total
1	7	2	4	3	51
	6	5	2	4 5	
2	8	1	-		20
	2	4		5	
3		2	7	9 1	48
		5	5	9 3	
4			2		25
			2	20	
Total	42	23	79		144

No. of Observations:

		P			
		1	2	3	
1		4	2	5	11
2	£	4	0	1	5
3		2	2	6	10
4		0	3	1	4
		10	7	13	30

To get P (adjusted for α) we set up:

11	0	0	0	4	2	5	51	} I
5	0	0	0	4	0	1	20	
10	0	0	0	2	2	6	48	
4	0	0	0	0	3	1	25	
				10	0	0	42	} II
				7	0	0	23	
				13	0	0	79	

Doolittle part I:

11 0 0 0	4	2	5	51
1 0 0 0	4/11	2/11	5/11	51/11
5 0 0	4	0	1	20
1 0 0	4/5	0	1/5	4
10 0	2	2	6	48
1 0	2/10	2/10	6/10	48/10
4	0	3	1	25
1	0	3/4	1/4	25/4

At this point we eliminate \hat{a} using (4) and (5)

$$c_{11} = n_{1.} - \sum \frac{a_1 n_{11}^2}{n_{1.}} = 10$$

$$- \left(\frac{4}{11} \times 4 + \frac{4}{5} \times 4 + \frac{2}{10} \times 2 \right) = \frac{2720}{550}$$

$$c_{12} = - \sum \frac{a n_{11} n_{12}}{n_{1.}} = - \left(\frac{4}{11} \times 2 + \frac{2}{10} \times 2 \right) = - \frac{620}{550}$$

$$g_1 = Y_{.1.} - \sum \frac{a n_{11} Y_{1..}}{n_{1.}}$$

$$= 42 - \left(\frac{4}{11} \times 51 + \frac{4}{5} \times 20 + 0 \times 25 \right) = - \frac{2360}{1100}$$

and so on and so forth.

$$C\hat{P} = 9$$

2720	620	2100	2360
550	550	550	1100
620	4385	3145	16.085
550	1100	1100	1100
2100	3145	7345	18445
550	1100	1100	1100

Now add 1/3, convert to decimal and Doolittle we get:
 (for convenience rounded to 2 decimals from hereon) :

5.27	- .79	-3.48	- 2.14	1	0	0
	4.31	-2.52	-14.62	0	1	0
		7.01	16.76	0	0	1
5.27	- .79	-3.48	- 2.14	1	0	0
1	- .15	- .66	- .40	.18	0	0
	4.20	-3.04	-14.94	.15	1	0
	1	- .72	- 3.55	.03	.23	0
		2.49	4.49	.76	.72	1
		1	1.80	.30	.29	.44

$$b_1 = 1.80(1) = 1.80$$

$$b_2 = 1.8(.72) + (-3.55)(1.0) = -2.24$$

$$b_3 = 1.8(.76) + (-3.55)(.15) + (-.40)(1) = .44$$

Then get the inverted matrix:

.43	.25	.30
.25	.44	.29
.30	.29	.40

Any row or column of the above should add to 1. Finally subtract $1/3$ from each element of the foregoing matrix we get the inverse:

.09	-.07	-.02
-.07	.11	-.04
-.02	-.04	.06

Any row or column of the above should add to zero.

$$\begin{aligned} \Sigma \text{CPID} = SS^P(\text{adj: for } \alpha) &= -2.14(-.40) + (-14.94)(-.355) \\ &\quad + 4.49(1.80) \\ &= 62.1644 \end{aligned}$$

Analyses of Variance

P (adj for α)

Source	d.f.	S.S.
Total	29	400.8000
α (unadj)	3	11.9045
P (adj for α)	2	$\Sigma \text{CPID} = 62.1644$
αP	4	by subtr. = 203.9811
Within Cells	20	122.7500

α (adj for P)

Source	d.f.	S.S.
Total	29	400.8000
P (unadj)	2	40.8483
α (adj for P)	3	α (unadj) + P (adj)
αP	4	- P (unadj) = 33.2206
Within Cells	20	203.9811
		122.7500

